

Ejercicio 5

Sean $X_1, X_2, \dots, X_n \text{ iid} \sim F$ Encontrar los estimadores de máxima verosimilitud para los siguientes parámetros y compararlos con los respectivos estimadores por el método de los momentos:

1. p si la distribución es $\text{Ber}(p)$
2. λ si la distribución es $\mathcal{P}(\lambda)$
3. p si la distribución es $\text{Geo}(p)$
4. μ y σ^2 si la distribución es $N(\mu, \sigma^2)$
5. a y b si la distribución es $\mathcal{U}[a, b]$.

Si $X_1, \dots, X_n \text{ iid} \sim \mathcal{U}[a, b]$ (x_1, \dots, x_n) valores observados

$$f_X(x) = \begin{cases} \frac{1}{b-a} & \text{si } x \in [a, b] \\ 0 & \text{si no} \end{cases}$$

estimadores por máxima verosimilitud para a, b

* función de verosimilitud

$$L(a, b) = \prod_{i=1}^n f_X(x_i) = \begin{cases} \frac{1}{(b-a)^n} & \text{si todos los } x_i \in [a, b] \\ 0 & \text{si no} \end{cases}$$

* tomamos logaritmo de $L(a, b)$

$$\ell(a, b) = \log(L(a, b)) = \log \left(\frac{1}{(b-a)^n} \right) = \log((b-a)^{-n}) \\ \uparrow \\ x_i \in [a, b] \qquad \qquad \qquad = -n \log(b-a)$$

$$\frac{\partial \ell}{\partial a}(a, b) = -n \frac{-1}{b-a} = \frac{n}{b-a} > 0 \text{ no tiene puntos críticos}$$

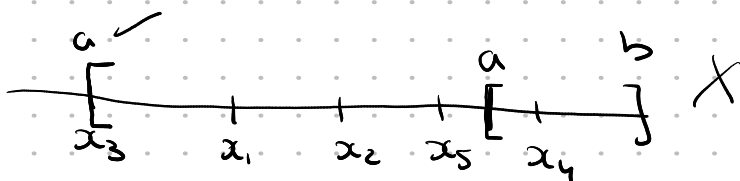
$$\frac{\partial \ell}{\partial b}(a, b) = -n \frac{1}{b-a} = -\frac{n}{b-a} < 0 \text{ no tiene puntos críticos}$$

$$L(a, b) = \begin{cases} \frac{1}{(b-a)^n} & \text{si } x_i \in [a, b] \\ 0 & \text{si no} \end{cases}$$

$\frac{1}{(b-a)^n}$ es más grande mientras más chico sea $b-a$
 $b > a$

→ queremos que b sea lo más chico posible y que a sea lo más grande posible

$U[a, b]$ (x_1, \dots, x_n) valores observados



los valores observados tienen que pertenecer al intervalo $[a, b]$

→ queremos que a sea lo más grande posible y además $a \leq x_i$ para todo $i \in \{1, \dots, n\}$

$$\hat{a} = \min \{x_1, \dots, x_n\}$$

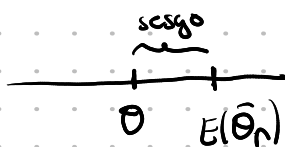
→ queremos que b sea lo más chico posible y además $x_i \leq b$ para todo $i \in \{1, \dots, n\}$

$$\hat{b} = \max \{x_1, \dots, x_n\}$$

Sesgo de un estimador

$\hat{\theta}_n$ estimador de θ

$$\begin{aligned} \text{sesgo}(\hat{\theta}_n) &= E(\hat{\theta}_n - \theta) \\ &= E(\hat{\theta}_n) - E(\theta) \\ &= E(\hat{\theta}_n) - \theta \end{aligned}$$



Decimos que $\hat{\theta}_n$ es insesgado si $\text{sesgo}(\hat{\theta}_n) = 0$ o equivalentemente si $E(\hat{\theta}_n) = \theta$.

ejemplo: $\sigma_n^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2$ y $s_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$ son estimadores de

la varianza σ_n^2 no es insesgado

s_n^2 es insesgado

Ejercicio 9

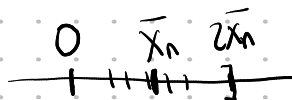
Sean X_1, \dots, X_n iid $\sim U[0, \theta]$. Interesa estimar el valor de θ .

1. Hallar el estimador de θ por el método de los momentos.
2. Estudiar su sesgo, varianza y error cuadrático medio.
3. Demostrar que el estimador de máxima verosimilitud de θ es X_n^* , el máximo de los valores muestrales.

$$X_1, \dots, X_n \text{ iid } \sim U[0, \theta]$$

① estimador de θ por el método de los momentos

$$\bar{X}_n \xrightarrow{cs} E(X_1) = \frac{\theta}{2}$$



$$\bar{X}_n \xrightarrow{cs} \frac{\theta}{2}$$

estimador de θ : $\hat{\theta}_n = 2\bar{X}_n$

② * sesgo de $\hat{\theta}_n$

$$\theta = E(\hat{\theta}_n)$$

$$\text{sesgo}(\hat{\theta}_n) = E(\hat{\theta}_n) - \theta$$

$$E(\hat{\theta}_n) = E(2\bar{X}_n) = 2E\left(\frac{X_1 + \dots + X_n}{n}\right)$$

$$= \frac{2}{n} E(X_1 + \dots + X_n)$$

$$= \frac{2}{n} \left(\underbrace{E(X_1)}_{\theta/2} + \dots + \underbrace{E(X_n)}_{\theta/2} \right)$$

$$= \frac{2}{n} \cdot n \cdot \frac{\theta}{2}$$

$$= \theta$$

$$\text{sesgo}(\hat{\theta}_n) = E(\hat{\theta}_n) - \theta = \theta - \theta = 0$$

$\Rightarrow \hat{\theta}_n$ es un estimador insesgado de θ

* $\text{Var}(\hat{\theta}_n)$

$$\text{Var}(\hat{\theta}_n) = \text{Var}(2\bar{X}_n) = 4 \text{Var}(\bar{X}_n)$$

$$= 4 \text{Var}\left(\frac{X_1 + \dots + X_n}{n}\right)$$

$$= \frac{4}{n^2} \text{Var}(X_1 + \dots + X_n)$$

$$X \sim U(a, b)$$

$$\text{Var}(X) = \frac{(b-a)^2}{12}$$

$$= \frac{4}{n^2} \left(\underbrace{\text{Var}(X_1)}_{=\frac{\theta^2}{12}} + \dots + \underbrace{\text{Var}(X_n)}_{\frac{\theta^2}{12}} \right)$$

porque las X_i son independientes

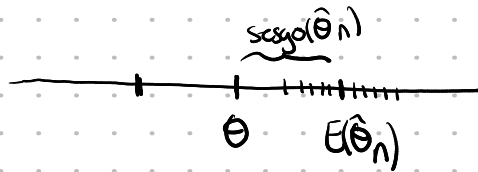
$$= \frac{4}{n^2} n \frac{\theta^2}{12}$$

$$= \frac{\theta^2}{3n}$$

Error cuadrático medio

$\hat{\theta}_n$ estimador de θ

$$\text{ECM}(\hat{\theta}_n) = E((\hat{\theta}_n - \theta)^2) = \text{Var}(\hat{\theta}_n) + \text{sesgo}(\hat{\theta}_n)^2$$



$$\text{ECM}(\hat{\theta}_n) = E((\hat{\theta}_n - \theta)^2)$$

$$\text{Var}(\hat{\theta}_n) = E(\hat{\theta}_n^2) - E(\hat{\theta}_n)^2$$

$$= E(\hat{\theta}_n^2 - 2\hat{\theta}_n\theta + \theta^2)$$

$$= E(\hat{\theta}_n^2) - 2\theta E(\hat{\theta}_n) + \theta^2$$

$$= E(\hat{\theta}_n^2) - E(\hat{\theta}_n)^2 + E(\hat{\theta}_n)^2 - 2\theta E(\hat{\theta}_n) + \theta^2$$

$$= \text{Var}(\hat{\theta}_n) + \underbrace{(E(\hat{\theta}_n) - \theta)^2}_{\text{sesgo}(\hat{\theta}_n)}$$

$$= \text{Var}(\hat{\theta}_n) + \text{sesgo}(\hat{\theta}_n)^2$$

⌞

$$\times \text{ECM}(\hat{\theta}_n)$$

$$\begin{aligned} \text{ECM}(\hat{\theta}_n) &= \text{Var}(\hat{\theta}_n) + \underbrace{\text{sesgo}(\hat{\theta}_n)^2}_{=0} \\ &= \frac{\theta^2}{3n} \end{aligned}$$

Teorema central del límite

X_1, \dots, X_n MAS de X

$$E(X) = \mu, \text{Var}(X) = \sigma^2$$

Consideramos \bar{X}_n

$$* E(\bar{X}_n) = \mu$$

$$\begin{aligned} E(\bar{X}_n) &= E\left(\frac{X_1 + \dots + X_n}{n}\right) = \frac{1}{n} E(X_1 + \dots + X_n) \\ &= \frac{1}{n} (E(X_1) + \dots + E(X_n)) \\ &= \frac{1}{n} \underbrace{E(X) = \mu}_{E(X) = \mu} + \dots + \underbrace{E(X) = \mu}_{E(X) = \mu} \\ &= \frac{1}{n} n \mu \end{aligned}$$

$$* \text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$$

$$\begin{aligned} \text{Var}(\bar{X}_n) &= \text{Var}\left(\frac{X_1 + \dots + X_n}{n}\right) = \frac{1}{n^2} \text{Var}(X_1 + \dots + X_n) \\ &= \frac{1}{n^2} (\underbrace{\text{Var}(X_1)}_{\text{Var}(X) = \sigma^2} + \dots + \underbrace{\text{Var}(X_n)}_{\text{Var}(X) = \sigma^2}) \quad \left. \begin{array}{l} \text{porque las } X_i \\ \text{son independientes} \end{array} \right\} \\ &= \frac{1}{n^2} n \sigma^2 \\ &= \frac{1}{n} \sigma^2 \end{aligned}$$

TCL : para n grande \bar{X}_n tiene distribución normal

$$\bar{X}_n \underset{\text{aprox}}{\sim} \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\frac{(\bar{X}_n - \mu)}{\sqrt{\frac{\sigma^2}{n}}} = \frac{(\bar{X}_n - \mu)}{\sigma} \sqrt{n} \underset{\text{aprox}}{\sim} \mathcal{N}(0, 1)$$

Ejercicio 2

Los resistores de cierto tipo de tienen resistencias que en promedio son de $\mu = 200$ ohms, con una desviación estándar de $\sigma = 10$ ohms. Se toman (al azar) 25 de estos resistores y se conectan (en forma independiente) en un circuito.

1. Calcular la probabilidad (aproximada) de que la resistencia promedio de los 25 resistores este entre 199 y 202 ohms.
2. Calcular la probabilidad (aproximada) de que la resistencia total de los 25 resistores no sea mayor que 5100 ohms.

X = valor de la resistencia de un resistor

$$E(X) = 200$$

$$\text{Var}(X) = \sigma^2 = 100$$

X_1, \dots, X_{25} MAS de X

$$\textcircled{1} P(199 \leq \bar{X}_{25} \leq 202)$$

$$E(\bar{X}_{25}) = E(X) = 200$$

$$\text{Var}(\bar{X}_{25}) = \frac{\text{Var}(X)}{25} = \frac{100}{25} = 4 = 2^2$$

por el TCL: $\bar{X}_{25} \sim_{\text{aprox}} \mathcal{N}(200, 2^2)$

$$P(199 \leq \bar{X}_{25} \leq 202) = P\left(\frac{199-200}{2} \leq \frac{\bar{X}_{25}-200}{2} \leq \frac{202-200}{2}\right)$$

$$= P\left(-\frac{1}{2} \leq \underbrace{\frac{\bar{X}_{25}-200}{2}}_{\sim_{\text{aprox}} \mathcal{N}(0,1)} \leq 1\right)$$

$$\approx \phi(1) - \phi\left(-\frac{1}{2}\right)$$

$$= \phi(1) - (1 - \phi\left(\frac{1}{2}\right))$$

$$= \phi(1) + \phi\left(\frac{1}{2}\right) - 1$$

$$= 0,8413 + 0,6915 - 1$$

$$= 0,5328$$

| z | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 | 0.05 |
|-------|--------|--------|--------|--------|--------|--------|
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 | 0.5199 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 | 0.5596 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 | 0.5987 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 | 0.6368 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 | 0.6736 |
| → 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 | 0.7088 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 | 0.7422 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7703 | 0.7734 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 | 0.8023 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 | 0.8289 |
| → 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 | 0.8531 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 | 0.8749 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 | 0.8944 |

$$\textcircled{2} P(X_1 + X_2 + \dots + X_{25} \leq 5100)$$

Forma 1:

$$P(X_1 + \dots + X_{25} \leq 5100) = P\left(\frac{X_1 + \dots + X_{25}}{25} \leq \frac{5100}{25}\right)$$

$$\bar{X}_{25} \sim_{\text{aprox}} \mathcal{N}(200, 2^2)$$

$$= P(\bar{X}_{25} \leq 204)$$

$$= P\left(\frac{\bar{X}_{25} - 200}{2} \leq \frac{204 - 200}{2}\right)$$

$$= P\left(\frac{\bar{X}_{25} - 200}{2} \leq 2\right)$$

$$\sim_{\text{aprox}} \mathcal{N}(0, 1)$$

$$\approx \Phi(2)$$

$$= 0,9772$$

| z | 0.00 | 0.01 | 0.02 | 0.03 | 0.04 |
|-----|--------|--------|--------|--------|--------|
| 0.0 | 0.5000 | 0.5040 | 0.5080 | 0.5120 | 0.5160 |
| 0.1 | 0.5398 | 0.5438 | 0.5478 | 0.5517 | 0.5557 |
| 0.2 | 0.5793 | 0.5832 | 0.5871 | 0.5910 | 0.5948 |
| 0.3 | 0.6179 | 0.6217 | 0.6255 | 0.6293 | 0.6331 |
| 0.4 | 0.6554 | 0.6591 | 0.6628 | 0.6664 | 0.6700 |
| 0.5 | 0.6915 | 0.6950 | 0.6985 | 0.7019 | 0.7054 |
| 0.6 | 0.7257 | 0.7291 | 0.7324 | 0.7357 | 0.7389 |
| 0.7 | 0.7580 | 0.7611 | 0.7642 | 0.7673 | 0.7703 |
| 0.8 | 0.7881 | 0.7910 | 0.7939 | 0.7967 | 0.7995 |
| 0.9 | 0.8159 | 0.8186 | 0.8212 | 0.8238 | 0.8264 |
| 1.0 | 0.8413 | 0.8438 | 0.8461 | 0.8485 | 0.8508 |
| 1.1 | 0.8643 | 0.8665 | 0.8686 | 0.8708 | 0.8729 |
| 1.2 | 0.8849 | 0.8869 | 0.8888 | 0.8907 | 0.8925 |
| 1.3 | 0.9032 | 0.9049 | 0.9066 | 0.9082 | 0.9099 |
| 1.4 | 0.9192 | 0.9207 | 0.9222 | 0.9236 | 0.9251 |
| 1.5 | 0.9332 | 0.9345 | 0.9357 | 0.9370 | 0.9382 |
| 1.6 | 0.9452 | 0.9463 | 0.9474 | 0.9484 | 0.9495 |
| 1.7 | 0.9554 | 0.9564 | 0.9573 | 0.9582 | 0.9591 |
| 1.8 | 0.9641 | 0.9649 | 0.9656 | 0.9664 | 0.9671 |
| 1.9 | 0.9713 | 0.9719 | 0.9726 | 0.9732 | 0.9738 |
| 2.0 | 0.9772 | 0.9778 | 0.9783 | 0.9788 | 0.9793 |
| 2.1 | 0.9821 | 0.9826 | 0.9830 | 0.9834 | 0.9838 |
| 2.2 | 0.9861 | 0.9864 | 0.9868 | 0.9871 | 0.9875 |
| 2.3 | 0.9893 | 0.9896 | 0.9898 | 0.9901 | 0.9904 |

Forma 2:

$\Gamma X_1, \dots, X_n$ MAS X

$$E(X) = \mu, \text{Var}(X) = \sigma^2$$

$$E(X_1 + \dots + X_n) = \underbrace{E(X_1)}_{=E(X)} + \dots + \underbrace{E(X_n)}_{=E(X)} = n\mu$$

$$\left(E(X_1 + \dots + X_n) = E(n\bar{X}_n) = nE(\bar{X}_n) = n\mu \right)$$

$$\text{Var}(X_1 + \dots + X_n) = \text{Var}(X_1) + \dots + \text{Var}(X_n) = n\sigma^2$$

$$\left(\text{Var}(X_1 + \dots + X_n) = \text{Var}(n\bar{X}_n) = n^2 \text{Var}(\bar{X}_n) = n^2 \frac{\sigma^2}{n} = n\sigma^2 \right)$$

TCL: si n es grande $X_1 + \dots + X_n$ tiene distribución normal

$$X_1 + \dots + X_n \sim_{\text{aprox}} \mathcal{N}(n\mu, n\sigma^2)$$

$$\frac{(X_1 + \dots + X_n) - n\mu}{\sigma\sqrt{n}} \sim_{\text{aprox}} \mathcal{N}(0, 1)$$

L

X_1, \dots, X_{25} MAS de X

$$E(X) = 200$$

$$\text{Var}(X) = 100$$

$$E(X_1 + \dots + X_{25}) = 5000$$

$$\text{Var}(X_1 + \dots + X_{25}) = 2500 = 50^2$$

$$\text{TCL}: X_1 + \dots + X_{25} \sim_{\text{aprox}} \mathcal{N}(5000, 50^2)$$

$$P(X_1 + \dots + X_{25} \leq 5100) = P\left(\frac{X_1 + \dots + X_{25} - 5000}{50} \leq \frac{5100 - 5000}{50}\right)$$

$$= P\left(\underbrace{\frac{X_1 + \dots + X_{25} - 5000}{50}}_{\sim \mathcal{N}(0, 1)} \leq 2\right)$$

$$\approx \phi(2)$$

$$= 0,9772$$