

VECTOR OF RANDOM PROCESSES

Definitions

Let us consider an experiment whose result is represented by a vector of n random processes $\mathbf{X}(t) = \{X_1(t) X_2(t) \dots X_n(t)\}^T$ (for example the components of the seismic motion at the base of n piers of a viaduct, the wind velocities registered by n anemometers, the dynamic response of a n -D.O.F. system). The vector $\mathbf{X}(t)$ is also called a n -variate random process.

Let us consider the values $x_i^{(j)}(t_i)$ ($i = 1, 2, \dots, n; j = 1, 2, \dots$) assumed by the sample functions $x_i^{(j)}(t)$ of $X_i(t)$ for $t = t_i$. The set of these values constitutes n random variables $X_i = X(t_i)$ ($i = 1, 2, \dots, n$). They are characterised by the joint density function of order n $p_{X_1 X_2 \dots X_n}(x_1, t_1; x_2, t_2; \dots x_n, t_n) = p_{\mathbf{X}}(\mathbf{x}, \mathbf{t})$, where $\mathbf{x} = \{x_1 x_2 \dots x_n\}^T$ is the vector of the state variables and $\mathbf{t} = \{t_1 t_2 \dots t_n\}^T$ is the vector of the times in correspondence of which the random variables are extracted.

If $\mathbf{X}(t)$ is a n -variate normal random process, its complete probabilistic definition calls for the knowledge of only the joint density functions of the second order of all the possible couples of random variables $X_i(t_i), X_j(t_j)$ ($i, j = 1, 2, \dots, n$): $p_{X_i X_j}(x_i, t_i; x_j, t_j), \forall i, j = 1 \dots n$ and $\forall t_i, t_j \in \mathbf{R}$.

Stationary processes

If $\mathbf{X}(t)$ is a (weakly) stationary random process, the joint density functions of the second order of all the possible couples of random variables $X_i(t_i), X_j(t_j)$ ($i, j = 1, 2, \dots, n$) do not depend on the instants t_i, t_j but only on the time lag $\tau = t_j - t_i$: $p_{X_i X_j}(x_i, t_i; x_j, t_j) = p_{X_i X_j}(x_i, x_j; \tau), \forall i, j = 1 \dots n$ and $\forall \tau \in \mathbf{R}$.

Let us define as mean vector $\boldsymbol{\mu}_{\mathbf{X}}$ of the stationary process $\mathbf{X}(t)$, a vector whose components $\mu_{X_i}(t)$ are the means of the processes $X_i(t)$ ($i = 1, \dots, n$):

$$\boldsymbol{\mu}_{\mathbf{X}} = E[\mathbf{X}(t)] = \{\mu_{X_1} \mu_{X_2} \dots \mu_{X_n}\}^T \quad (44)$$

Let us define as correlation matrix and covariance matrix of $\mathbf{X}(t)$, respectively, the matrices whose i, j -th terms are the functions $R_{X_i X_j}(\tau)$ and $C_{X_i X_j}(\tau)$:

$$\mathbf{R}_{\mathbf{X}}(\tau) = E[\mathbf{X}(t) \mathbf{X}^T(t + \tau)] = \begin{bmatrix} R_{X_1 X_1}(\tau) & R_{X_1 X_2}(\tau) & \dots & R_{X_1 X_n}(\tau) \\ R_{X_2 X_1}(\tau) & R_{X_2 X_2}(\tau) & \dots & R_{X_2 X_n}(\tau) \\ \vdots & \vdots & \ddots & \vdots \\ R_{X_n X_1}(\tau) & R_{X_n X_2}(\tau) & \dots & R_{X_n X_n}(\tau) \end{bmatrix} \quad (45)$$

$$\mathbf{C}_{\mathbf{X}}(\tau) = E[\{\mathbf{X}(t) - \boldsymbol{\mu}_{\mathbf{X}}\} \{\mathbf{X}(t + \tau) - \boldsymbol{\mu}_{\mathbf{X}}\}^T] = \begin{bmatrix} C_{X_1 X_1}(\tau) & C_{X_1 X_2}(\tau) & \dots & C_{X_1 X_n}(\tau) \\ C_{X_2 X_1}(\tau) & C_{X_2 X_2}(\tau) & \dots & C_{X_2 X_n}(\tau) \\ \vdots & \vdots & \ddots & \vdots \\ C_{X_n X_1}(\tau) & C_{X_n X_2}(\tau) & \dots & C_{X_n X_n}(\tau) \end{bmatrix} \quad (46)$$

The on-diagonal terms are, respectively, the auto-correlation functions and the auto-covariance functions of each process; the off-diagonal terms are, respectively, the cross-correlation functions and the cross-covariance functions of all the possible couples of the processes.

Eqs. (45) and (46) involve:

$$\mathbf{R}_X(\tau) = \mathbf{C}_X(\tau) + \boldsymbol{\mu}_X \boldsymbol{\mu}_X^T \quad (47)$$

Due to Eqs. (16) and (21):

$$\mathbf{R}_X(\tau) = \mathbf{R}_X^T(-\tau) \quad (48)$$

$$\mathbf{C}_X(\tau) = \mathbf{C}_X^T(-\tau) \quad (49)$$

Thus, the matrices \mathbf{R}_X and \mathbf{C}_X are symmetric for $\tau = 0$:

$$\mathbf{R}_X(0) = \mathbf{R}_X^T(0) \quad (50)$$

$$\mathbf{C}_X(0) = \mathbf{C}_X^T(0) \quad (51)$$

It can be shown that \mathbf{R}_X and \mathbf{C}_X are semi-positive definite for any value of τ ($\mathbf{f}_R = \mathbf{u}^T \mathbf{R}_X \mathbf{u} \geq 0$, $\mathbf{f}_C = \mathbf{v}^T \mathbf{C}_X \mathbf{v} \geq 0$ for $\forall \mathbf{u}, \mathbf{v} \neq \mathbf{0}$).

Let us define as power spectral density matrix of $\mathbf{X}(t)$ the matrix whose i,j -th term is the function $S_{X_i X_j}(\omega)$:

$$\mathbf{S}_X(\omega) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \mathbf{C}_X(\tau) e^{-i\omega\tau} d\tau = \begin{bmatrix} S_{X_1 X_1}(\omega) & S_{X_1 X_2}(\omega) & \dots & S_{X_1 X_n}(\omega) \\ S_{X_2 X_1}(\omega) & S_{X_2 X_2}(\omega) & \dots & S_{X_2 X_n}(\omega) \\ \vdots & \vdots & \vdots & \vdots \\ S_{X_n X_1}(\omega) & S_{X_n X_2}(\omega) & \dots & S_{X_n X_n}(\omega) \end{bmatrix} \quad (52)$$

$$\mathbf{C}_X(\tau) = \int_{-\infty}^{\infty} \mathbf{S}_X(\omega) e^{i\omega\tau} d\omega = \begin{bmatrix} C_{X_1 X_1}(\tau) & C_{X_1 X_2}(\tau) & \dots & C_{X_1 X_n}(\tau) \\ C_{X_2 X_1}(\tau) & C_{X_2 X_2}(\tau) & \dots & C_{X_2 X_n}(\tau) \\ \vdots & \vdots & \vdots & \vdots \\ C_{X_n X_1}(\tau) & C_{X_n X_2}(\tau) & \dots & C_{X_n X_n}(\tau) \end{bmatrix} \quad (53)$$

The on-diagonal terms of the spectral matrix are the power spectral densities of each process; the off-diagonal terms are the cross-power spectral densities of all the possible couples of the processes.

Due to Eq. (37) $\mathbf{S}_X(\omega)$ is a Hermitian matrix. It can be proved that it is also semi-positive definite for any value of ω . From Eq. (53) it results:

$$\Sigma = C_X(0) = \int_{-\infty}^{\infty} S_X(\omega) d\omega = \begin{bmatrix} \phi_{X_1}^2 & C_{X_1X_2}(0) & \dots & C_{X_1X_n}(0) \\ C_{X_2X_1}(0) & \phi_{X_2}^2 & \dots & C_{X_2X_n}(0) \\ \vdots & \vdots & \ddots & \vdots \\ C_{X_nX_1}(0) & C_{X_nX_2}(0) & \dots & \phi_{X_n}^2 \end{bmatrix} \quad (54)$$

The on-diagonal terms are the mean square values of each random process; the off-diagonal terms are the cross-covariance functions (i.e. the cross-correlation functions for a zero mean process) of all the possible couples of processes, at the time lag $\tau = 0$.

n-variate normal process

A random stationary process $\mathbf{X}(t)$ has a normal n-variate distribution if all the couples of the random processes that constitute $\mathbf{X}(t)$ have a joint density function of order 2 given by the Eq. (24).

A particular linear transformation

Let $\mathbf{Y}(t)$ be a n-variate process linked with the n-variate process $\mathbf{X}(t)$ by the relationship:

$$\mathbf{Y}(t) = \mathbf{A}\mathbf{X}(t)$$

It can be shown that:

$$\boldsymbol{\mu}_Y = \mathbf{A}\boldsymbol{\mu}_X$$

$$\mathbf{C}_Y(\tau) = \mathbf{A}\mathbf{C}_X(\tau)\mathbf{A}^T$$

$$\mathbf{R}_Y(\tau) = \mathbf{A}\mathbf{R}_X(\tau)\mathbf{A}^T$$

$$\mathbf{S}_Y(\omega) = \mathbf{A}\mathbf{S}_X(\omega)\mathbf{A}^T$$