# Applications of Information Theory in Image Processing

**1. Review of Information Theory and Lossless Source Coding** 

# Notation

- A : discrete (usually finite) alphabet;  $\alpha = |A|$  : size of A (when finite)
- $A^n$ : set of strings of length *n* over *A*;  $A^*$ : set of finite strings over *A*
- $\lambda$  : empty string
- $x_1^n = x^n = x_1 x_2 \dots x_n$  : finite sequence over A
- $x_1^{\infty} = x^{\infty} = x_1 x_2 \dots x_t \dots$ : infinite sequence over *A*
- $x_i^j = x_i x_{i+1} \dots x_j$  : sub-sequence (*i* sometimes omitted if = 1)
- $p_X(x)$  or  $P_X(x)$  : Prob(X = x) probability mass function (PMF) on A (subscript X dropped if clear from context)
- $X \sim p(x) : X$  obeys PMF p(x)
- $E_p[F]$  : expectation of F w.r.t. PMF p (subscript and [] may be dropped)
- $\hat{p}_{x^n}(x)$  : empirical distribution obtained from  $x^n$
- $\log x$  : logarithm to base 2 of x, unless base otherwise specified
- $\ln x$  : natural logarithm of x
- H(X), H(p) : entropy of a r.v. X or PMF p, in bits (usually per-symbol)
- $H(X^n)$  : joint entropy of  $X_1, X_2, ..., X_n$  (unnormalized)
- $H_2(p) = -p \log p (1-p) \log(1-p)$ ,  $0 \le p \le 1$ : binary entropy function
- D(p||q) : relative entropy (information divergence) between PMFs p and q

### Coding in a communication/storage system



# **Information Theory**

#### □ Shannon, "A mathematical theory of communication," Bell Tech. Journal, 1948

- Theoretical foundations of source and channel coding
- Fundamental bounds and coding theorems in a probabilistic setting
  - in a nutshell: perfect communication in the presence of noise is possible as long as the *entropy rate* of the source is below the *channel capacity*
- Fundamental theorems essentially non-constructive: we've spent the last 70 years realizing Shannon's promised paradise in practice
  - very successful: enabled current digital revolution (multimedia, internet, wireless communication, mass storage, ...)
- Separation theorem: *source and channel coding can be done independently*

# **Source Coding**



# Source coding = Data compression $\Rightarrow$ efficient use of bandwidth/space

# **Data Compression**



- **\Box** Lossless compression: D = D' (the case of interest here)
- □ Lossy compression:  $D' \approx D$ ; D' is an *approximation* of D under some metric

# **Data Sources**



- □ Symbols  $x_i \in A = a$  countable (usually finite) alphabet.
- □ Probabilistic source:  $x_i$  are realizations of random variables  $X_i$ ;  $X_1^n$  obeys some probability distribution P on  $A^n$ .
- □ We are often interested in  $n \to \infty$ :  $X_1^{\infty}$  is a *random process*.
  - *stationary* (*time-invariant*):  $X_i^{\infty} = X_j^{\infty}$ , as random processes,  $\forall i, j \ge 1$
  - ergodic: time averages converge to ensemble averages
  - memoryless: X<sub>i</sub> are statistically independent
  - independent, identically distributed (i.i.d.): memoryless, and  $X_i \sim p_X \forall i$

# **Data Sources**



- □ Symbols  $x_i \in A = a$  countable (usually finite) alphabet.
- Individual sequence: x<sub>i</sub> are just symbols, not assumed to be a realization of a random process. We will talk about probability assignments, but they will be derived from the data under certain constraints, and with certain objectives.
- $\Box$  Here too, we will often be interested in asymptotic behavior as  $n \to \infty$ .

# **Statistics on Individual Sequences**

 $\Box$  Empirical distributions derived from an individual sequence  $x^n$ .

$$\hat{p}_{x^n}(a) = \frac{1}{n} \left| \{ i : 1 \le i \le n, x_i = a \} \right|, a \in A \qquad \text{Memoryless} \\ (\text{zero-th order})$$

- We can compute empirical statistics of *any order* (joint, conditional, etc.)
- Sequence probability according to its own empirical distribution

$$\widehat{P}_{x^n}(x^n) = \prod_{i=1}^n \widehat{p}_{x^n}(x_i)$$

- This is the *highest probability* assigned to the sequence by any distribution from the model class: *maximum likelihood (ML) probability*
- Example: *Bernoulli model* (binary, i.i.d.)

$$A = \{0, 1\}, \quad n_0 = \left| \{ i : x_i = 0\} \right|, \quad n_1 = n - n_0 = \left| \{ i : x_i = 1\} \right|$$

$$\hat{p}(0) = \frac{n_0}{n}, \ \hat{p}(1) = \frac{n_1}{n}: \ \hat{P}_{x^n}(x^n) = \hat{p}(0)^{n_0} \ \hat{p}(1)^{n_1} = \frac{n_0^{n_0} n_1^{n_1}}{n^n}$$

- Notice that if  $x^n$  is in fact the outcome of a random process, then its empirical distribution (and other statistics) are themselves random variables
  - e.g., expressions of the type  $P_{X^n}(|\hat{p}(a) p(a)| \ge \epsilon)$

# **Statistical Models for Data Sources: Finite Memory**



$$P(x_1^n) = \prod_{t=1}^n p(x_t | x_{t-k}^{t-1})$$

- Say  $|A| = \alpha$ . The  $\alpha \cdot \alpha^k$  numbers p(a|s),  $a \in A$ ,  $s \in A^k$ , together with the fixed initial state, completely define the source.
- In fact, there are only  $(\alpha 1) \cdot \alpha^k$  independent parameters; once we have p(a|s) for  $\alpha 1$  symbols a, the probability of the remaining symbol is fully determined.

## **Statistical Models for Data Sources: Finite Memory**

Sequence probability

$$P(x^{n}) = \prod_{t=1}^{n} p(x_{t}|x_{t-k}^{t-1})$$
  

$$rac{1}{k-th order Markov empirical distribution}$$
  

$$\hat{p}_{x^{n},k}(a|s) = \frac{\left|\left\{t: 1 \le t \le n, x_{t} = a, x_{t-k}^{t-1} = s\right\}\right|}{\left|\left\{t: 1 \le t \le n, x_{t-k}^{t-1} = s\right\}\right|}$$
  
number of times  
we see state s

 $\Box$  As before, we can ask what probability this distribution assigns to  $x^n$ 

$$\hat{P}_{x^{n},k}(x^{n}) = \prod_{t=1}^{n} \hat{p}_{x^{n},k}(x_{t}|x_{t-k}^{t-1})$$

• This is the ML probability of  $x^n$  for the class of k-th order Markov models.

# **Example: Binary finite memory source,** k = 2

$$p(0|00) = 1/4$$

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$$p(0|00) = 1/4$$

$$p(0|10) = 7/8$$

$$p(0|01) = 1/8$$

$$p(0|10) = 1/8$$

**Steady state:**  $\pi_{ab} \stackrel{\text{\tiny def}}{=} p_s(ab)$  stationary state probabilities

$$\pi_{ab} = \sum_{cd} \pi_{cd} P(ab|cd), \quad ab, cd \in \{00, 01, 10, 11\}$$
$$[\pi_{00}, \pi_{01}, \pi_{10}, \pi_{11}] = [\pi_{00}, \pi_{01}, \pi_{10}, \pi_{11}] \cdot \begin{bmatrix} \frac{1}{4} & \frac{3}{4} & 0 & 0\\ 0 & 0 & \frac{1}{8} & \frac{7}{8}\\ \frac{7}{8} & \frac{1}{8} & 0 & 0\\ 0 & 0 & \frac{3}{4} & \frac{1}{4} \end{bmatrix}$$

 $\Rightarrow [\pi_{00}, \pi_{01}, \pi_{10}, \pi_{11}] = \begin{bmatrix} \frac{7}{26} & \frac{6}{26} & \frac{7}{26} \end{bmatrix}, \qquad [p_{s}(0), p_{s}(1)] = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \end{bmatrix}$ 

 $p_{s}(b) = \sum_{cd} \pi_{cd} \cdot P(b|cd)$  stationary symbol probabilities

# **Statistical Models for Data Sources: FSM**

**Given State Machine (FSM)** 

- state space  $S = \{s_0, s_1, ..., s_{K-1}\}$
- initial state  $s_0$
- output probability  $p(a|s), a \in A, s \in S$
- state transition probability  $q(s|s', x), \quad s, s' \in S, \ x \in A$
- *unifilar*  $\Leftrightarrow$  deterministic transitions: *next-state function*  $f: S \times A \rightarrow S$
- every finite memory source is equivalent to a unifilar FSM with  $K \leq |A|^k$ , but in general, finite state  $\neq$  finite memory





# **Example: Binary FSM**



$$\Box \text{ Steady state: } \pi_i \stackrel{\text{def}}{=} p_{\text{stat}}(s_i)$$

$$[\pi_0 \pi_1 \pi_2] \begin{bmatrix} .9 & .1 & 0 \\ .5 & 0 & .5 \\ .1 & 0 & .9 \end{bmatrix} = [\pi_0 \pi_1 \pi_2]$$

$$\Rightarrow [\pi_0 \pi_1 \pi_2] = \begin{bmatrix} \frac{5}{8} & \frac{1}{16} & \frac{5}{16} \end{bmatrix}, [p_0 p_1] = \begin{bmatrix} \frac{5}{8} & \frac{3}{8} \end{bmatrix}$$

$$\xrightarrow{\text{stationary symbol probs.}}$$

$$p_{\text{stat}}(b) = \sum_i \pi_i \cdot p(b|s_i)$$

# **Statistical Models for Data Sources: Trees**

#### **Tree sources**

- finite memory  $\leq k$  (Markov)
- # of past symbols needed to determine the state might be
   < k for some states</li>



- by merging nodes from the full Markov tree, we get a model with a *smaller number of free parameters*
- the set of tree sources with unbalanced trees has *measure zero* in the space of Markov sources of any given order
- yet, tree source models have proven very useful in practice, and are associated with some of the best compression algorithms to date



 $s_0 \leftrightarrow 0, \quad s_1 \leftrightarrow 01, \quad s_2 \leftrightarrow 11$ 

this tree has an FSM representation (not always the case)

# Entropy

$$X \sim p(x)$$
:  $H(X) = -\sum_{x \in A} p(x) \log p(x)$ 

 $[\log = \log_2, \ 0 \log 0 \stackrel{\text{\tiny def}}{=} 0]$ 

 $H(X) = E_p[-\log p(X)]$ 

*entropy* of *X* (or of the PMF *p*), measured in *bits*.

- *H* measures the *uncertainty* or (the average of the) *self*-*information* of *X*.
- We also write H(p) : a random variable is not actually needed;
   p(·) could be an empirical distribution.

# **Entropy: example**

$$X \sim p(x)$$
:  $H(X) = -\sum_{x \in A} p(x) \log p(x)$  [ $0 \log 0 \stackrel{\text{def}}{=} 0$ ]

Example:  $A = \{0, 1\}$  P(0) = p, P(1) = 1 - p

 $H_2(p) = -p \log p - (1-p) \log(1-p)$  bin

binary entropy function

#### Main properties:

- $H_2(p) \ge 0$ ,  $H_2(p)$  is  $\cap$ -convex,  $0 \le p \le 1$
- $H_2(p) \rightarrow 0$  as  $p \rightarrow 0$  or 1, with slope  $\infty$

• 
$$H_2(p)$$
 is maximal at  $p = 0.5$ ,  $H_2(0.5) = 1$ 

⇒ the entropy of an unbiased coin is 1 bit



# Entropy (cont.)

□ For a general finite alphabet A, H(X) is maximal when X is *uniformly distributed*, i.e.,

$$X \sim p_u$$
, where  $p_{u(a)} = \frac{1}{|A|} \quad \forall a \in A$ 



Proof:

• Jensen's inequality:

if f is a  $\cup$ -convex function and Y a r.v., then  $Ef(Y) \ge f(EY)$ .

•  $-\log x$  is a  $\cup$ -convex function of x; set Y = 1/p(X) and  $f(Y) = -\log Y$ 



$$H(X) = E\left[\log\frac{1}{p(X)}\right] = E\left[\log Y\right] = -E\left[-\log Y\right] \le -(-\log EY)$$
$$= \log E\left[\frac{1}{p(X)}\right] = \log|A| = H(p_u)$$
Jensen

# **Joint Entropy**

□ The *joint entropy* of random variables  $(X, Y) \sim p(x, y)$  is defined as

$$\mathbf{H}(X,Y) = -\sum_{x,y} p(x,y) \log p(x,y)$$

• This can be extended to any number of random variables:  $H(X_1, X_2, ..., X_n)$ . Notation:

 $\mathbf{H}(X_1, X_2, \dots, X_n) = \text{joint entropy of } X_1, X_2, \dots, X_n \quad (0 \le \mathbf{H} \le n \log |A|)$ 

$$H(X_1, X_2, ..., X_n) = \frac{1}{n} \mathbf{H}(X_1, X_2, ..., X_n)$$

= normalized per-symbol entropy  $(0 \le H \le \log |A|)$ 

• If (X, Y) are statistically independent, then  $\mathbf{H}(X, Y) = \mathbf{H}(X) + \mathbf{H}(Y)$ .

# **Conditional Entropy**

$$\mathbf{H}(X,Y) = -\sum_{x,y} p(x,y) \log p(x,y)$$

The *conditional entropy* (of *Y* conditioned on *X*) is defined as

$$H(Y|X) = \sum_{x} p(x)H(Y|X = x)$$
$$= -\sum_{x} p(x) \sum_{y} p(y|x) \log p(y|x)$$
$$= -\sum_{x,y} p(x,y) \log p(y|x) = -E \log p(Y|X)$$

Chain rule

$$\mathbf{H}(X,Y) = H(X) + H(Y|X)$$

Conditioning reduces uncertainty

 $H(X|Y) \le H(X)$ 

• but  $H(X|Y = y) \ge H(X)$  is possible

# **Entropy Rates**

**□** Entropy rate of a random process

$$H(X_1^{\infty}) = \lim_{n \to \infty} \frac{1}{n} \mathbf{H}(X_1^n)$$

in bits/symbol, if the limit exists!

A related limit based on *conditional entropy* 

 $H^*(X_1^{\infty}) = \lim_{n \to \infty} H(X_n | X_{n-1}, X_{n-2}, \dots, X_1)$ 

in bits/symbol, if the limit exists!

Theorem: For a stationary random process, both limits exist, and

 $H^*(X_1^\infty)=H(X_1^\infty)$ 

# **Entropy rates (examples)**

$$\begin{array}{l} \blacksquare X_1, X_2, \dots \text{ i.i.d.:} \\ H(X_1^{\infty}) &= \lim_{n \to \infty} \frac{1}{n} \operatorname{H}(X_1, X_2, \dots, X_n) = \lim_{n \to \infty} \frac{1}{n} n H(X_1) = H(X_1) \\ \blacksquare X_1^{\infty} \quad \text{stationary } k \text{-th order Markov:} \\ H(X_1^{\infty}) &= H^*(X_1^{\infty}) & \text{theorem} \\ &= \lim_{n \to \infty} H(X_n | X_{n-1}, \dots, X_1) & \text{definition} \\ &= \lim_{n \to \infty} H(X_n | X_{n-1}, \dots, X_{n-k}) & \text{Markov} \\ &= H(X_{k+1} | X_k, \dots, X_1) & \text{stationary} \end{array}$$

The theorem provides a very useful tool to compute entropy rates for a broad family of source models

# **Entropy Rates - Example**



steady state:

 $[\pi_{00}, \pi_{01}, \pi_{10}, \pi_{11}] = \begin{bmatrix} \frac{7}{26} & \frac{6}{26} & \frac{7}{26} \end{bmatrix},$ 

 $[p_{s}(0), p_{s}(1)] = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \end{bmatrix}$ 

Markov process entropy

**Zero-order entropy**  $H\left(\frac{1}{2}\right) = 1$ 

$$H(X|S) = \sum_{ab} \pi_{ab} H(p(0|ab)) =$$
  
$$\frac{7}{26}H(\frac{1}{4}) + \frac{6}{26}H(\frac{1}{8}) + \frac{6}{26}H(\frac{7}{8}) + \frac{7}{26}H(\frac{3}{4}) \approx 0.688$$

# **Entropy Rates - Example**



Steady state:

$$\begin{bmatrix} \pi_0 & \pi_1 & \pi_2 \end{bmatrix} = \begin{bmatrix} \frac{5}{8} & \frac{1}{16} & \frac{5}{16} \end{bmatrix}, \quad \begin{bmatrix} p_0 & p_1 \end{bmatrix} = \begin{bmatrix} \frac{5}{8} & \frac{3}{8} \end{bmatrix}$$
state probs. symb. probs.

Zero-order entropy

H(0.375) = 0.954

■ Markov process entropy  $H(X | S) = \sum_{i=0}^{2} p(s_i) H(p(0 | s_i)) =$  $\frac{5}{8} H(0.9) + \frac{1}{16} H(0.5) + \frac{5}{16} H(0.1) \approx 0.502$ 

# Individual sequence - fitted with FSM model 00000011111111100000000001111111110 s<sub>0</sub> s<sub>1</sub> s<sub>2</sub>

**Empirical entropy:** 

$$\hat{p}(0 \mid s_0) = \frac{16}{19}, \ \hat{p}(0 \mid s_1) = \frac{1}{3}, \ \hat{p}(0 \mid s_2) = \frac{1}{9}, \quad \left[\hat{\pi}_0 \ \hat{\pi}_1 \ \hat{\pi}_2\right] = \left[\frac{19}{40} \ \frac{3}{40} \ \frac{18}{40}\right], \quad \hat{H}(x \mid S) = 0.594$$

# **Empirical entropy**

 $\Box$  Defined for a sequence  $x^n$ , relative to a class of models, as

$$\widehat{H}(x^n) = -\frac{1}{n}\log(\text{ML probability of } x^n)$$
 normalized,  
in bits/symbol

• Example: Bernoulli model. Recall

$$\hat{P}_{\chi^n}(x^n) = \hat{p}(0)^{n_0} \, \hat{p}(1)^{n_1} = \frac{n_0^{n_0} n_1^{n_1}}{n^n} = \left(\frac{n_0}{n}\right)^{n_0} \left(\frac{n_1}{n}\right)^{n_1} \qquad \text{using}_{n_0 + n_1 = n}$$

This is the ML probability of  $x^n$  relative to the class of Bernoulli models (zeroorder Markov).

• We have

$$\widehat{H}(x^n) = -\frac{1}{n}\log\widehat{P}_{x^n}(x^n) = -\frac{n_0}{n}\log\frac{n_0}{n} - \frac{n_1}{n}\log\frac{n_1}{n} = H_2(\widehat{p}(0))$$

in fact, "empirical entropy = entropy of empirical probability"
holds for most probability models we are interested in,
including Markov models of any order

# **Relative Entropy**

□ The relative entropy (or Kullback-Leibler distance, or information divergence) between two PMFs p(x) and q(x) is defined as

$$D(p||q) = \sum_{x} p(x)\log\frac{p(x)}{q(x)} = E_p\log\frac{p(x)}{q(x)}$$

• <u>Theorem</u>:  $D(p||q) \ge 0$ , with equality iff p = q.

Proof (using strict concavity of log, and Jensen's inequality):

$$-D(p||q) = \sum_{x} p(x) \log \frac{q(x)}{p(x)} \le \log \sum_{x} p(x) \frac{q(x)}{p(x)} = \log \sum_{x} q(x) \le 0$$

the summations are over values of x where  $p(x)q(x) \neq 0$ ; other terms contribute either 0 or  $\infty$  to D. Since log is strictly concave, equality holds iff  $\frac{p(x)}{q(x)} = 1 \forall x$ .

- *D* is not symmetric, and therefore not a distance in the metric sense.
- However, it is a very useful way to express 'proximity' of distributions.

in a sense, D(p||q) measures the inefficiency of assuming that the distribution is q when it is actually p